

Lecture 2 on data assimilation: The ensemble Kalman filter (the algebra of)

Marc Bocquet

With help from Alban Farchi, inspiration from Pavel Sakov

CEREA, joint lab École des Ponts ParisTech and EdF R&D, Université Paris-Est, France
Institut Pierre-Simon Laplace

(marc.bocquet@enpc.fr)



Synopsis of the course

- **Monday, October 28 10:30-12:30**

Lecture 1: Elementary principles of geophysical data assimilation. The Bayesian standpoint. Classical methods of data assimilation: 3D-Var, the Kalman filter, 4D-Var.

- **Tuesday, October 29, 10:30-12:30**

Lecture 2: The ensemble Kalman filter and its variants (focus on the algorithmic/mathematical aspects.)

- **Thursday, October 31, 10:30-12:30**

Lecture 3: Recent advances: hybrid and ensemble variational techniques. Discussion on what to expect from machine learning/deep learning.

Followed next week by:

- A course on data assimilation and stochastic filtering, particle filters by Dan Crisan (Imperial College, London)
- A course on big data and uncertainty quantification by Omar Ghattas (Uni. of Texas, Austin)

Outline

- 1 The ensemble Kalman filter
 - Reminder
 - Principles
 - Mathematical prerequisites
 - The ETKF
 - The EnSRF
 - The DEnKF
 - The serial EnKF
- 2 Making EnKF work: localisation and inflation
 - Diagnostic
 - Localisation
 - Inflation
 - Why they are necessary
 - Hybrid localisation
- 3 References

Sequential Bayesian estimation

- Recall our HMM given by the dynamical model and observation model:

$$\mathbf{x}_k = \mathcal{M}_{k:k-1}(\mathbf{x}_{k-1}, \boldsymbol{\lambda}) + \boldsymbol{\eta}_k, \quad \mathbf{y}_k = \mathcal{H}_k(\mathbf{x}_k) + \boldsymbol{\epsilon}_k.$$

- The model and the observational errors, $\boldsymbol{\eta}_k, \boldsymbol{\epsilon}_k : k = 1, \dots, K$ are assumed to be **uncorrelated in time, mutually independent**, and they follow the pdfs $p_{\boldsymbol{\eta}}$ and $p_{\boldsymbol{\epsilon}}$.

Formal sequential Bayesian solution

- An **analysis** step, in which the conditional pdf $p(\mathbf{x}_k | \mathbf{y}_{k:0})$ is updated using the latest observation vector, \mathbf{y}_k ,

$$p(\mathbf{x}_k | \mathbf{y}_{k:0}) \propto p_{\boldsymbol{\eta}}(\mathbf{y}_k - \mathcal{H}_k(\mathbf{x}_k)) p(\mathbf{x}_k | \mathbf{y}_{k-1:0}),$$

- which alternates with a **forecast** step which propagates this pdf, using the Chapman-Kolmogorov equation, forward in time until the new observation batch:

$$p(\mathbf{x}_{k+1} | \mathbf{y}_{k:0}) = \int d\mathbf{x} p_{\boldsymbol{\eta}}(\mathbf{x}_k - \mathcal{M}_{k:k-1}(\mathbf{x}_{k-1})) p(\mathbf{x}_k | \mathbf{y}_{k:0}).$$

Sequential Bayesian estimation: the Kalman filter

► Even though these equations are well suited for **sequential** DA with **chaotic** models, they are still impractical to solve. However, the **Kalman filter** solves them exactly under the assumptions of **linearity of the models** and **Gaussianity of the statistics**.

► **Analysis** step:

$$\mathbf{x}_k^a = \mathbf{x}_k^f + \mathbf{K}_k \left(\mathbf{y}_k - \mathbf{H}_k \mathbf{x}_k^f \right),$$

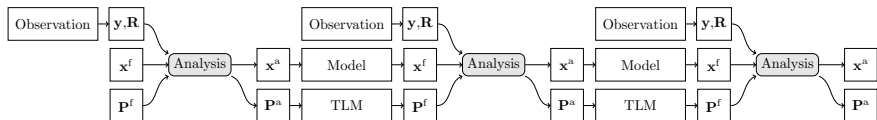
$$\mathbf{K}_k = \mathbf{P}_k^f \mathbf{H}_k^\top \left(\mathbf{R}_k + \mathbf{H}_k \mathbf{P}_k^f \mathbf{H}_k^\top \right)^{-1},$$

$$\mathbf{P}_k^a = \left(\mathbf{I}_x - \mathbf{K}_k \mathbf{H}_k \right) \mathbf{P}_k^f.$$

► **Forecast** step:

$$\mathbf{x}_{k+1}^f = \mathbf{M}_{k+1:k} \mathbf{x}_k^a,$$

$$\mathbf{P}_{k+1}^f = \mathbf{M}_{k+1:k} \mathbf{P}_k^a \mathbf{M}_{k+1:k}^\top + \mathbf{Q}_{k+1}.$$



The extended Kalman filter

- ▶ As seen in lecture 1, the Kalman filter can be extended to handle **nonlinear** models:

$$\begin{aligned}\mathbf{x}_{k+1}^f &= \mathcal{M}_{k+1:k}(\mathbf{x}_k^a), \\ \mathbf{P}_{k+1}^f &= \mathbf{M}_{k+1:k} \mathbf{P}_k^a \mathbf{M}_{k+1:k}^\top + \mathbf{Q}_{k+1},\end{aligned}$$

where $\mathbf{M}_{k+1:k}$ is the tangent linear model (linearisation at \mathbf{x}_k^a) of $\mathcal{M}_{k+1:k}$.

- ▶ **Drawbacks 1 & 2:** Extremely costly for large geophysical models: storage space (storage of \mathbf{P}^f) and computations ($\mathbf{M}_{k+1:k} \mathbf{P}_k^f \mathbf{M}_{k+1:k}^\top$ requires $2N_x$ integrations of the model).
- ▶ **Drawback 3:** The model linearisation in the error covariances is an approximation.
- ▶ **Solutions:** The reduced-rank / ensemble Kalman filters.

The ensemble Kalman filter

- The idea [Evensen 1994; Houtekamer and Mitchell 1998] is to make the KF work in high dimensions and replace \mathbf{P} (\mathbf{P}^a and \mathbf{P}^f) with an ensemble of states $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_{N_e}$. The moments of the error could theoretically be approximated by **the sample/empirical moments**:

$$\bar{\mathbf{x}}^f = \frac{1}{N_e} \sum_{i=1}^{N_e} \mathbf{x}_i^f, \quad \mathbf{P}^f \approx \frac{1}{N_e - 1} \sum_{i=1}^{N_e} (\mathbf{x}_i^f - \bar{\mathbf{x}}^f) (\mathbf{x}_i^f - \bar{\mathbf{x}}^f)^\top.$$

- Define the normalised **anomaly** or **perturbation** matrix $\in \mathbb{R}^{N_x \times N_e}$

$$[\mathbf{X}_f]_i = \frac{\mathbf{x}_i^f - \bar{\mathbf{x}}^f}{\sqrt{N_e - 1}} \quad \Rightarrow \quad \mathbf{P}^f \approx \mathbf{X}_f \mathbf{X}_f^\top.$$

Likewise

$$\bar{\mathbf{x}}^a = \frac{1}{N_e} \sum_{i=1}^{N_e} \mathbf{x}_i^a, \quad \mathbf{P}^a \approx \mathbf{X}_a \mathbf{X}_a^\top \quad \text{where} \quad [\mathbf{X}_a]_i = \frac{\mathbf{x}_i^a - \bar{\mathbf{x}}^a}{\sqrt{N_e - 1}}.$$

The ensemble Kalman filter: Ansatz and mean update

- ▶ An educated guess would suggest, for $i = 1 \dots N_e$:

$$\mathbf{x}_i^a = \mathbf{x}_i^f + \mathbf{K} \left(\mathbf{y} - \mathbf{H}\mathbf{x}_i^f \right).$$

but the **correct** answer is actually

$$\mathbf{x}_i^a = \mathbf{x}_i^f + \mathbf{K} \left(\mathbf{y} + \boldsymbol{\epsilon}_i - \mathbf{H}\mathbf{x}_i^f \right).$$

where $\boldsymbol{\epsilon}_i$ is a **stochastic noise** sampled from $\mathcal{N}(\mathbf{0}, \mathbf{R})$, for each member.

- ▶ **Checking the mean:** on average, and summing over the ensemble members:

$$\bar{\mathbf{x}}^a = \bar{\mathbf{x}}^f + \mathbf{K} \left(\mathbf{y} - \mathbf{H}\bar{\mathbf{x}}^f \right),$$

which is the same as the Kalman filter's mean update.

The ensemble Kalman filter: perturbations update

► **Checking the ensemble update:** on average, does it mimic the Kalman filter?

We define

$$\bar{\epsilon} = \frac{1}{N_e} \sum_{i=1}^{N_e} \epsilon_i, \quad \Theta = \frac{1}{\sqrt{N_e - 1}} [\epsilon_1 - \bar{\epsilon} \quad \epsilon_2 - \bar{\epsilon} \quad \cdots \quad \epsilon_{N_e} - \bar{\epsilon}].$$

The perturbations update then reads (ensemble minus the mean):

$$\mathbf{X}_a = (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{X}_f + \mathbf{K}\Theta,$$

which yields the empirical analysis error covariances:

$$\mathbf{P}^a = (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{P}^f(\mathbf{I}_x - \mathbf{K}\mathbf{H})^\top + \mathbf{K}\Theta\Theta^\top\mathbf{K}^\top + (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{X}_f\Theta^\top\mathbf{K}^\top + \mathbf{K}\Theta\mathbf{X}_f^\top(\mathbf{I}_x - \mathbf{K}\mathbf{H})^\top,$$

whose average on Θ is

$$\mathbb{E}[\mathbf{P}^a] = (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{P}^f(\mathbf{I}_x - \mathbf{K}\mathbf{H})^\top + \mathbf{K}\mathbf{R}\mathbf{K}^\top = (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{P}^f.$$

The last identity is valid if \mathbf{K} is the (optimal) Kalman gain.

► In the absence of the observation stochastic noise, the posterior error statistics would be incorrect!

The ensemble Kalman filter: forecast

- Kalman gain representations:

Empirical: denoting $\mathbf{Y}_f = \mathbf{H}\mathbf{X}_f + \Theta$, we have $\mathbf{K} = \mathbf{X}_f \mathbf{Y}_f^\top (\mathbf{Y}_f \mathbf{Y}_f^\top)^{-1}$

Deterministic: denoting $\mathbf{Y}_f = \mathbf{H}\mathbf{X}_f$, we have $\mathbf{K} = \mathbf{X}_f \mathbf{Y}_f^\top (\mathbf{R} + \mathbf{Y}_f \mathbf{Y}_f^\top)^{-1}$

- Forecast step: The ensemble is propagated using the full nonlinear model

$$\mathbf{x}_{i,k+1}^f = \mathcal{M}_{k+1:k}(\mathbf{x}_{i,k}^a),$$

whereas the extended Kalman filter uses the tangent linear model.

- Numerically costly (N_e propagations) but
 - the forecast scheme is **embarrassingly parallel**,
 - no need to derive the tangent linear model of the full model.

The ensemble Kalman filter: surrogate for \mathbf{H}

- Instead of estimating $\mathbf{P}^f \mathbf{H}^\top = \mathbf{X}_f \mathbf{Y}_f^\top$ and $\mathbf{H} \mathbf{P}^f \mathbf{H}^\top = \mathbf{Y}_f \mathbf{Y}_f^\top$ in the Kalman gain, we can use the ensemble:

$$\bar{\mathbf{y}}^f = \frac{1}{N_e} \sum_{i=1}^{N_e} \mathcal{H}(\mathbf{x}_i^f),$$

$$\mathbf{P}^f \mathbf{H}^\top = \frac{1}{N_e - 1} \sum_{i=1}^{N_e} (\mathbf{x}_i^f - \bar{\mathbf{x}}^f) [\mathcal{H}(\mathbf{x}_i^f) - \bar{\mathbf{y}}^f]^\top,$$

$$\mathbf{H} \mathbf{P}^f \mathbf{H}^\top = \frac{1}{N_e - 1} \sum_{i=1}^{N_e} [\mathcal{H}(\mathbf{x}_i^f) - \bar{\mathbf{y}}^f] [\mathcal{H}(\mathbf{x}_i^f) - \bar{\mathbf{y}}^f]^\top.$$

These approximations rely on the key assumption:

$$[\mathbf{Y}_f]_i = \mathbf{H} (\mathbf{x}_i^f - \bar{\mathbf{x}}^f) \approx \mathcal{H}(\mathbf{x}_i^f) - \bar{\mathbf{y}}^f.$$

- This is sometimes called the **secant method** (alternative to finite-differences).

The ensemble Kalman filter: What's nice about it?

The ensemble forecast has a complexity of N_e model runs

Yes, it is far better than the extended Kalman filter and game-changing.
But there will be a heavy tribute for this.

The ensemble forecast uses the nonlinear model in place of the tangent linear model

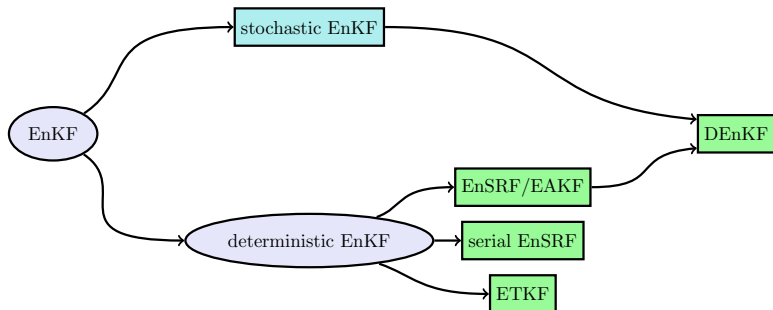
Yes, it's nice and better from a Bayesian standpoint.
But not as critical as it was originally sold. In that respect, the EnKF is outperformed by the **iterative** ensemble Kalman filter and smoother (→ lecture 3).

It emulates the tangent linear of the observation model

Definitely a good point and at the origin of nonlinear **EnVar** techniques (→ lecture 3).

The ensemble Kalman filter: a bunch of methods

- ▶ Two main flavors of EnKFs: **stochastic** and **deterministic**, but many variants.



- ▶ But several significant precursors and alternatives: reduced-rank square-root Kalman filter, SEEK, SEIK, unscented Kalman filter, etc.

Key algebraic identities

- ▶ **Sherman-Morrison-Woodbury** (SMW) identity (**A** and **C** invertible):

$$(\mathbf{A} + \mathbf{UCV})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1}\mathbf{U}(\mathbf{C}^{-1} + \mathbf{VA}^{-1}\mathbf{U})^{-1}\mathbf{VA}^{-1}.$$

- ▶ Typical applications:

- Analysis error covariances:

$$\mathbf{P}^a = (\mathbf{B}^{-1} + \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H})^{-1} = \mathbf{B} - \mathbf{B}\mathbf{H}^\top (\mathbf{R} + \mathbf{H}\mathbf{B}\mathbf{H}^\top)^{-1} \mathbf{H}\mathbf{B}.$$

- Kalman gain:

$$\mathbf{K} = \mathbf{B}\mathbf{H}^\top (\mathbf{R} + \mathbf{H}\mathbf{B}\mathbf{H}^\top)^{-1} = (\mathbf{B}^{-1} + \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H})^{-1} \mathbf{H}^\top \mathbf{R}^{-1}.$$

Key algebraic identities

- **Matrix shift lemma** (SML): Let \mathbf{A} and \mathbf{B} two matrices of compatible dimensions and $x \mapsto f(x)$ be a function defined on the spectra of \mathbf{AB} and \mathbf{BA} , then :

$$\mathbf{A}f(\mathbf{BA}) = f(\mathbf{AB})\mathbf{A}.$$

→ Proof in [Higham 2008].

- Typical application, $\mathbf{A} \in \mathbb{R}^{N_x \times N_y}$ and $\mathbf{B} \in \mathbb{R}^{N_y \times N_x}$ are positive semi-definite:

$$\mathbf{A}(\mathbf{I}_y + \mathbf{BA})^{-1} = (\mathbf{I}_x + \mathbf{AB})^{-1}\mathbf{A}.$$

Key algebraic identities

► Let f be a function such that $f(0) = 1$, and which is analytic in a connected domain \mathcal{D} of contour \mathcal{C} in the complex plane \mathbb{C} which encloses the eigenvalues of both \mathbf{AB} and \mathbf{BA} . Define $g(x) = (f(x) - 1)/x$. Then

$$f(\mathbf{AB}) = \mathbf{I} + \mathbf{A}g(\mathbf{BA})\mathbf{B}.$$

→ Proof in [Higham 2008].

► Application: let us assume that the eigenvalues of \mathbf{AB} and \mathbf{BA} have a non-negative real part, then

$$(\mathbf{I}_x + \mathbf{AB})^{-\frac{1}{2}} = \mathbf{I}_x - \mathbf{A} \left(\mathbf{I}_y + \mathbf{BA} + [\mathbf{I}_y + \mathbf{BA}]^{\frac{1}{2}} \right)^{-1} \mathbf{B},$$

where we chose $f(x) = (1+x)^{-\frac{1}{2}}$ and $g(x) = -(1+x+\sqrt{1+x})^{-1}$.

→ Proof in [Bocquet and Farchi 2019].

Deterministic Kalman filters and matrix square root definition

- The **deterministic EnKFs** avoid the introduction of the stochastic perturbations by updating the anomaly matrix \mathbf{X}_f in

$$\mathbf{P}^f = \mathbf{X}_f \mathbf{X}_f^T,$$

rather than updating \mathbf{P}^f .

- In the following, \mathbf{X}_f is called a **factor** of \mathbf{P}^f , not a “square root” of \mathbf{P}^f as sometimes seen in geophysical data assimilation literature. This would clash with the mathematical definition of a square root matrix.

- Let \mathbf{M} be a diagonalisable matrix with non-negative eigenvalues, i.e. $\mathbf{M} = \mathbf{G} \mathbf{D} \mathbf{G}^{-1}$, where \mathbf{G} is an invertible matrix and \mathbf{D} is the diagonal matrix containing the non-negative eigenvalues of \mathbf{M} . Then the **square root** of \mathbf{M} is

$$\mathbf{M}^{\frac{1}{2}} = \mathbf{G} \mathbf{D}^{\frac{1}{2}} \mathbf{G}^{-1},$$

where $\mathbf{D}^{\frac{1}{2}}$ is the diagonal matrix with the square root of the eigenvalues of \mathbf{M} .

- Note that \mathbf{M} does not have to be symmetric.

The ensemble transform Kalman filter: mean update

- ▶ One of the variant (ETKF, [Hunt et al. 2007] on an idea by [Bishop, Etherton, et al. 2001]) operates the linear algebra **in the space of the perturbations**, or **ensemble subspace**:

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{X}_f \mathbf{w}^a.$$

- ▶ Inserting this decomposition into the Kalman state update equation:

$$\mathbf{x}^f + \mathbf{X}_f \mathbf{w}^a = \mathbf{x}^f + \mathbf{X}_f \mathbf{X}_f^\top \mathbf{H}^\top \left(\mathbf{H} \mathbf{X}_f \mathbf{X}_f^\top \mathbf{H}^\top + \mathbf{R} \right)^{-1} \delta, \quad \text{where} \quad \delta = \mathbf{y} - \mathcal{H}(\mathbf{x}^f),$$

which suggests

$$\mathbf{w}^a \equiv \mathbf{X}_f^\top \mathbf{H}^\top \left(\mathbf{H} \mathbf{X}_f \mathbf{X}_f^\top \mathbf{H}^\top + \mathbf{R} \right)^{-1} \delta = \mathbf{Y}_f^\top \left(\mathbf{Y}_f \mathbf{Y}_f^\top + \mathbf{R} \right)^{-1} \delta.$$

- ▶ Using the SMW identity, we finally obtain:

$$\mathbf{w}^a = \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{-1} \mathbf{Y}_f^\top \mathbf{R}^{-1} \delta.$$

The ensemble transform Kalman filter: perturbations update

- From the analysis error covariance matrix of the Kalman filter, let us infer what the analysis anomaly matrix could be:

$$\begin{aligned}
 \mathbf{P}^a &= (\mathbf{I}_x - \mathbf{K}\mathbf{H})\mathbf{P}^f \\
 &\approx \left(\mathbf{I}_x - \mathbf{X}_f \mathbf{Y}_f^\top (\mathbf{Y}_f \mathbf{Y}_f^\top + \mathbf{R})^{-1} \mathbf{H} \right) \mathbf{X}_f \mathbf{X}_f^\top \\
 &\approx \mathbf{X}_f \left(\mathbf{I}_e - \mathbf{Y}_f^\top (\mathbf{Y}_f \mathbf{Y}_f^\top + \mathbf{R})^{-1} \mathbf{Y}_f \right) \mathbf{X}_f^\top,
 \end{aligned}$$

which suggests to choose the following factor:

$$\mathbf{X}_a = \mathbf{X}_f \left(\mathbf{I}_e - \mathbf{Y}_f^\top (\mathbf{Y}_f \mathbf{Y}_f^\top + \mathbf{R})^{-1} \mathbf{Y}_f \right)^{1/2}.$$

The ensemble transform Kalman filter: perturbations update

- ▶ This expression can be simplified into

$$\begin{aligned}
 \mathbf{X}_a &= \mathbf{X}_f \left(\mathbf{I}_e - \mathbf{Y}_f^\top (\mathbf{Y}_f \mathbf{Y}_f^\top + \mathbf{R})^{-1} \mathbf{Y}_f \right)^{1/2} \\
 &\stackrel{\text{SMW}}{=} \mathbf{X}_f \left(\mathbf{I}_e - \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{-1} \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{1/2} \\
 &= \mathbf{X}_f \left[\left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{-1} \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f - \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right) \right]^{1/2} \\
 &= \mathbf{X}_f \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{-1/2}.
 \end{aligned}$$

- ▶ We conclude

$$\mathbf{X}_a = \mathbf{X}_f \mathbf{T}, \quad \text{with} \quad \mathbf{T} = \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{Y}_f \right)^{-1/2}.$$

- ▶ Now, we can build the posterior ensemble as

$$i = 1, \dots, N_e: \quad \mathbf{x}_i^a = \bar{\mathbf{x}}^a + \sqrt{N_e - 1} \mathbf{X}_f [\mathbf{T}]_i = \bar{\mathbf{x}}^f + \mathbf{X}_f \left(\mathbf{w}^a + \sqrt{N_e - 1} [\mathbf{T}]_i \right).$$

The ensemble transform Kalman filter: rotation matrix

- A more general anomaly update is

$$\mathbf{X}_a = \mathbf{X}_f \mathbf{T} \mathbf{U}, \quad \text{where} \quad \mathbf{U} \in O(N_e).$$

- It is important to require:

$$\mathbf{U} \mathbf{1} = \mathbf{1}, \quad \text{where} \quad \mathbf{1} = [1, \dots, 1]^T \in \mathbb{R}^{N_e}.$$

This ensures that the updated ensemble is **centred on \mathbf{x}^a** [Livings et al. 2008; Sakov and Oke 2008b]. Indeed, we have

$$\mathbf{X}_a \mathbf{1} = \mathbf{X}_f \mathbf{T} \mathbf{U} \mathbf{1} = \mathbf{X}_f \mathbf{T} \mathbf{1} = \mathbf{X}_f \mathbf{1} = \mathbf{0},$$

and

$$\frac{1}{N_e} \sum_{i=1}^{N_e} \mathbf{x}_i^a = \bar{\mathbf{x}}^a + \frac{\sqrt{N_e - 1}}{N_e} \mathbf{X}_a \mathbf{1} = \bar{\mathbf{x}}^a.$$

- $\mathbf{U} = \mathbf{I}_e$ minimises the distance between \mathbf{X}_a and \mathbf{X}_f [Ott et al. 2004]. However, choosing random \mathbf{U} may make the update more Gaussian and hence be more consistent with the EnKF assumptions [Lawson and Hansen 2004; Sakov and Oke 2008b].

- $\mathbf{U} = \mathbf{I}_e$ in the following for the sake of simplicity.

The ensemble square-root Kalman filter (EnSRF)

- ▶ This is a variant of the deterministic EnKF where the update is carried out **in state space**, rather than in **ensemble subspace** as for the ETKF.
- ▶ **Mean update**: same as all the other EnKFs.
- ▶ **Perturbation update** [Sakov and Bertino 2011]:

$$\begin{aligned}
 \mathbf{X}_a &= \mathbf{X}_f \left(\mathbf{I}_e + \mathbf{Y}_f^\top \mathbf{R}^{-1} \mathbf{H} \mathbf{X}_f \right)^{-\frac{1}{2}} \\
 &\stackrel{\text{SML}}{=} \left(\mathbf{I}_x + \mathbf{X}_f \mathbf{X}_f^\top \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{-\frac{1}{2}} \mathbf{X}_f \\
 &= \left(\mathbf{I}_x + \mathbf{P}^f \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{-\frac{1}{2}} \mathbf{X}_f.
 \end{aligned}$$

Very elegant formula though not practical!

Note that $\mathbf{I}_e + \mathbf{P}^f \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H}$ is in general **not symmetric** but it is **diagonalisable** with positive spectrum hence, it has a square root, which is unique.

- ▶ The EnSRF is algebraically equivalent and shares the left transform update with the adjustment ensemble Kalman filter (EAKF) [J. L. Anderson 2001].

DEnKF: the *deterministic* ensemble Kalman filter

- Reformulation of the perturbation update on the left:

We use $(\mathbf{I}_x + \mathbf{A}\mathbf{B})^{-\frac{1}{2}} = \mathbf{I}_x - \mathbf{A} \left(\mathbf{I}_y + \mathbf{B}\mathbf{A} + [\mathbf{I}_y + \mathbf{B}\mathbf{A}]^{\frac{1}{2}} \right)^{-1} \mathbf{B}$ with $\mathbf{A} = \mathbf{P}^f \mathbf{H}^\top$ and $\mathbf{B} = \mathbf{R}^{-1} \mathbf{H}$ and we obtain:

$$\begin{aligned} \mathbf{X}_a &= \left(\mathbf{I}_e + \mathbf{P}^f \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{\frac{1}{2}} \mathbf{X}_f \\ &= \left\{ \mathbf{I}_x - \mathbf{P}^f \mathbf{H}^\top \left(\mathbf{R} + \mathbf{H} \mathbf{P}^f \mathbf{H}^\top + \mathbf{R} \left[\mathbf{I}_y + \mathbf{R}^{-1} \mathbf{H} \mathbf{P}^f \mathbf{H}^\top \right]^{\frac{1}{2}} \right)^{-1} \mathbf{H} \right\} \mathbf{X}_f. \end{aligned}$$

- **Effective gain** in a deterministic setup:

Mimicking the stochastic EnKF, the effective gain for the updated perturbations (not the mean!) is

$$\tilde{\mathbf{K}} = \mathbf{P}^f \mathbf{H}^\top \left(\mathbf{R} + \mathbf{H} \mathbf{P}^f \mathbf{H}^\top + \mathbf{R} \left[\mathbf{I}_y + \mathbf{R}^{-1} \mathbf{H} \mathbf{P}^f \mathbf{H}^\top \right]^{\frac{1}{2}} \right)^{-1},$$

as shown by [Whitaker and Hamill 2002] following [Andrews 1968], [Farchi and Bocquet 2019]. This can be reformulated as

$$\tilde{\mathbf{K}} = \mathbf{K} \left\{ \mathbf{I}_y + \left(\mathbf{I}_y + \mathbf{H} \mathbf{P}^f \mathbf{H}^\top \mathbf{R}^{-1} \right)^{-\frac{1}{2}} \right\}^{-1}.$$

DEnKF: the deterministic ensemble Kalman filter

- ▶ An approximation of the EnSRF that mimics the update of the stochastic EnKF.
- ▶ **Mean update:** same as all the other EnKFs.
- ▶ In the weak assimilation regime, we have:

$$\left\{ \mathbf{I}_y + \left(\mathbf{I}_y + \mathbf{H}\mathbf{P}^f\mathbf{H}^\top\mathbf{R}^{-1} \right)^{-\frac{1}{2}} \right\}^{-1} \approx \frac{1}{2}\mathbf{I}_y.$$

which suggests that the effective gain matrix can be approximated as

$$\hat{\mathbf{K}} = \frac{1}{2}\mathbf{K},$$

i.e.

$$\mathbf{x}_a \approx \left(\mathbf{I}_x - \frac{1}{2}\mathbf{K}\mathbf{H} \right) \mathbf{x}_f.$$

- ▶ Avoids the need to compute the square root \rightarrow very similar to the stochastic EnKF (but deterministic).

DEnKF: the deterministic ensemble Kalman filter

- Why this filter is **robust**:

$$\begin{aligned}
 \widehat{\mathbf{P}}^a &= \widehat{\mathbf{X}}_a \widehat{\mathbf{X}}_a^\top = \left(\mathbf{I}_x - \frac{1}{2} \mathbf{K} \mathbf{H} \right) \mathbf{X}_f \mathbf{X}_f^\top \left(\mathbf{I}_x - \frac{1}{2} \mathbf{H}^\top \mathbf{K}^\top \right) \\
 &= \mathbf{P}^f - \frac{1}{2} \mathbf{K} \mathbf{H} \mathbf{P}^f - \frac{1}{2} \mathbf{P}^f \mathbf{H}^\top \mathbf{K}^\top + \frac{1}{4} \mathbf{K} \mathbf{H} \mathbf{P}^f \mathbf{H}^\top \mathbf{K}^\top \\
 &= (\mathbf{I}_x - \mathbf{K} \mathbf{H}) \mathbf{P}^f + \frac{1}{4} \mathbf{K} \mathbf{H} \mathbf{P}^f \mathbf{H}^\top \mathbf{K}^\top \\
 &\geq (\mathbf{I}_x - \mathbf{K} \mathbf{H}) \mathbf{P}^f = \mathbf{P}^a,
 \end{aligned}$$

i.e. the analysis error covariance matrix of the DEnKF ($\widehat{\mathbf{P}}^a$) is bounded by the exact one:

$$\widehat{\mathbf{P}}^a \geq \mathbf{P}^a.$$

- **Ensemble update**: In summary,

$$\mathbf{x}_i^a = \mathbf{x}_i^f + \mathbf{K} \left[\mathbf{y} - \mathcal{H} \left(\frac{\mathbf{x}_i^f + \bar{\mathbf{x}}^f}{2} \right) \right].$$

This nicely mimics the stochastic EnKF – the update can be carried out in parallel.

- Used in several intermediate and operational systems.

Serial EnKF

► Alternatively, the observations can be assimilated **one at a time**.

- Drawback: can lead to **suboptimality** whenever an approximation is introduced.
- Advantage: simple (especially the Potter scheme) and localisation is effective and elegant in this framework.

→ Used in the NCAR DART DA suite, and in most of J. L. Anderson's papers.

► **Mean update:**

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{K}(y - \mathfrak{h}(\mathbf{x}^f)) \quad \mathbf{K} = \mathbf{P}^f \mathbf{h}^\top / (r + \mathbf{h} \mathbf{P}^f \mathbf{h}^\top).$$

► **Perturbation update:**

$$\tilde{\mathbf{K}} = \frac{\mathbf{K}}{1 + 1/\sqrt{1 + r^{-1} \mathbf{h} \mathbf{P}^f \mathbf{h}^\top}}.$$

Outline

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 - Principles
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 - The ETKF
 - The EnSRF
 - The DEnKF
 - The serial EnKF
- 2 Making EnKF work: localisation and inflation
 - Diagnostic
 - Localisation
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 - Why they are necessary
 - Hybrid localisation
- 3 References

Remedies to make EnKF work in high dimension

- ▶ Limited number N_e of anomalies: the sample covariance matrix is **highly rank-deficient**.

- ▶ If \mathbf{B} is the true covariance matrix and \mathbf{P}^e is the (N_e -member) sample covariance matrix which approximates \mathbf{B} , then:

$$\mathbb{E} \left([\mathbf{P}^e - \mathbf{B}]_{ij}^2 \right) = \frac{1}{N_e - 1} \left([\mathbf{B}]_{ij}^2 + [\mathbf{B}]_{ii} [\mathbf{B}]_{jj} \right). \quad (1)$$

- In most geophysical systems, $[\mathbf{B}]_{ij}$ vanish exponentially with $|i - j| \rightarrow \infty$. The $[\mathbf{B}]_{ii}$ are the variances and remain finite, so that

$$\mathbb{E} \left([\mathbf{P}^e - \mathbf{B}]_{ij}^2 \right) \underset{|i-j| \rightarrow \infty}{\sim} \frac{1}{N_e - 1} [\mathbf{B}]_{ii} [\mathbf{B}]_{jj}. \quad (2)$$

- ▶ Since $[\mathbf{B}]_{ij}$ vanish exponentially with the distance, we expect $\mathbb{E} \left([\mathbf{P}^e - \mathbf{B}]_{ij}^2 \right)$ to also vanish exponentially with the distance. Hence with N_e finite, the sample covariance $[\mathbf{P}^e]_{ij}$ is potentially a bad approximation especially for large distances $|i - j|$.

- ▶ The errors of such an approximation are usually referred to as **sampling errors**.

Localisation

► **Covariance localisation** seeks to regularise the sample covariance to mitigate the rank-deficiency of \mathbf{P}^e and the appearance of spurious correlations.

► Solution: compute the **Schur product** of \mathbf{P}^e with a well chosen smooth **correlation matrix** ρ , that has exponentially vanishing correlations for distant parts.

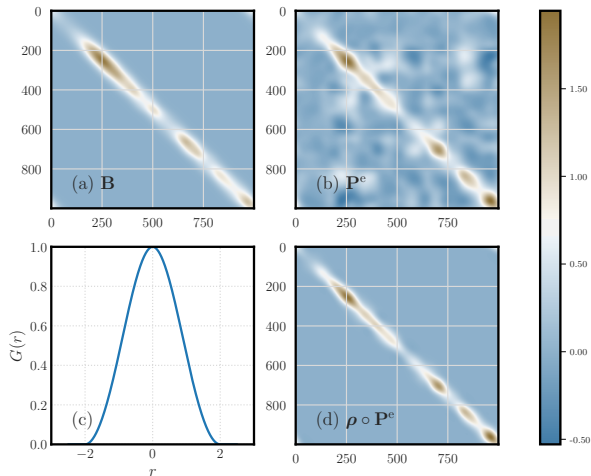
The Schur product of ρ and \mathbf{B} is defined by (**tapering** of covariances)

$$[\rho \circ \mathbf{P}^e]_{ij} = [\rho]_{ij} [\mathbf{P}^e]_{ij}. \quad (3)$$

Applicable only if the long-range error correlations are negligible.

► The Schur product theorem ensures that this product is **positive semi-definite**, a proper covariance matrix. For sufficiently regular ρ , $\rho \circ \mathbf{P}^e$ turns out to be **full-rank**.

Covariance localisation with the Gaspari-Cohn function



Panel (a): True covariance matrix. Panel (b): Sample covariance matrix.

Panel (c): Gaspari-Cohn based correlation matrix used for covariance localisation.

Panel (d): Tapered covariance matrix.

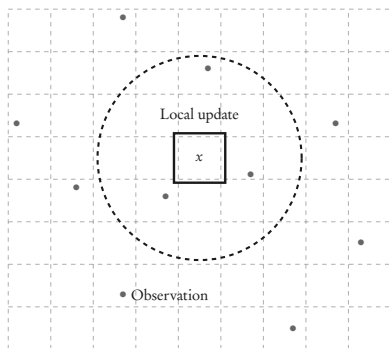
Domain localisation

- **Domain localisation**: divide & conquer.

The DA analysis is performed in parallel in **local domains**. The outcomes of these analyses are later sewed together.

Applicable only if the long-range error correlations are negligible.

Elegant but nor suited for the assimilation of non-local observations such as radiances.



- Both localisation schemes have successfully been applied to the EnKF [Hamill et al. 2001; Houtekamer and Mitchell 2001; Evensen 2003; Hunt et al. 2007].

Inflation

- ▶ Localisation addresses the rank-deficiency issue, but **sampling errors** are not entirely removed in the process: long EnKF runs may still diverge!
- ▶ Ad hoc means to counteract sampling errors is to **inflate the error covariance matrix** by a multiplicative factor $\lambda^2 \geq 1$:

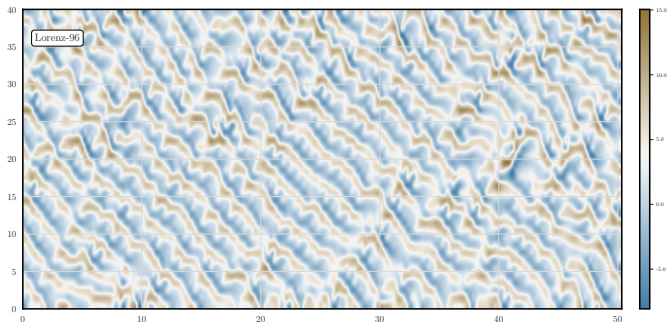
$$\mathbf{P}^e \longrightarrow \lambda^2 \mathbf{P}^e, \quad (4)$$

or, alternatively,

$$\mathbf{x}_{[n]} \longrightarrow \bar{\mathbf{x}} + \lambda \left(\mathbf{x}_{[n]} - \bar{\mathbf{x}} \right). \quad (5)$$

- ▶ Inflation can also come in an **additive form**: $\mathbf{x}_{[n]} \longrightarrow \mathbf{x}_{[n]} + \boldsymbol{\epsilon}_{[n]}$.
- ▶ Note that inflation is not only used to cure sampling errors, but is also often used to counteract **model error** impact.
- ▶ As a drawback, inflation often needs to be tuned, which is numerically costly. Hence, **adaptive** schemes have been developed to make the task more automatic [El Gharamti 2018; Raanes et al. 2019].

Nonlinear chaotic models: the Lorenz-96 low-order model



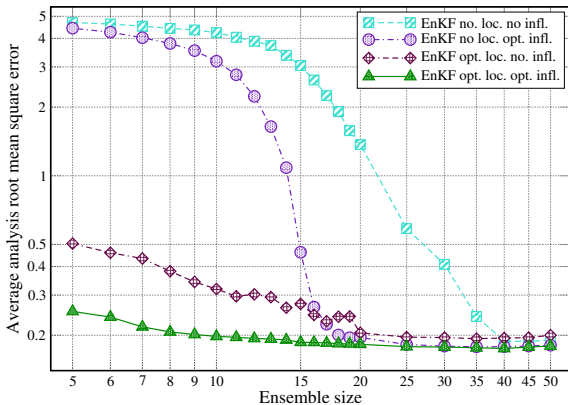
- It represents a mid-latitude zonal circle of the global atmosphere.
- Set of $N_x = 40$ ordinary differential equations [Lorenz and Emanuel 1998]:

$$\frac{dx_n}{dt} = (x_{n+1} - x_{n-2})x_{n-1} - x_n + F, \quad (6)$$

where $F = 8$, and the boundary is cyclic.

- Conservative system except for a forcing term F and a dissipation term $-x_n$.
- Chaotic dynamics, 13 positive and 1 neutral Lyapunov exponents, a doubling time of about 0.42 time units.

Illustration with the Lorenz-96 model



- Performance of the EnKF in the absence/presence of inflation/localisation.

The local ensemble transform Kalman filter (LETKF)

- ▶ Since the ETKF update is carried out in ensemble subspace, **only domain localisation** can be used. Hence an ETKF update is performed for each local domain.
- ▶ Advantages: The scheme is simple. Local ETKF updates are computed **in parallel**.
- ▶ Drawback: it is not possible to assimilate **nonlocal** observations such as radiances, without drastic approximations.

- ▶ Updating N_x variables with an ETKF could be seen as a formidable task. However, (i) the updates are parallel (ii) each local update operates on a reduce observation vector which drastically reduces the local numerical cost.

Mean update of the local EnKF (except for the LETKF)

- The mean analysis in the local EnKF is carried out using the Kalman gain matrix

$$\mathbf{K} = \mathbf{B}\mathbf{H}^\top \left(\mathbf{R} + \mathbf{H}\mathbf{B}\mathbf{H}^\top \right)^{-1}, \quad (7)$$

where \mathbf{H} is the observation operator (or tangent-linear thereof), and where the regularised

$$\mathbf{B} = \rho \circ \mathbf{P}^e \quad (8)$$

is used in place of the sample \mathbf{P}^e .

→ numerically very costly!

- Usually applied in **observation space** whenever the observations can be seen as point-wise, i.e. local. Then $\mathbf{B}\mathbf{H}^\top \approx \rho_{xy} \circ (\mathbf{P}^e \mathbf{H}^\top)$ and $\mathbf{H}\mathbf{B}\mathbf{H}^\top \approx \rho_{yy} \circ (\mathbf{H}\mathbf{P}^e \mathbf{H}^\top)$ where ρ_{xy} represents ρ acting in the cross product of the state and observations spaces and ρ_{yy} represents ρ acting in the observations space. As a result:

$$\mathbf{K} \approx \rho_{xy} \circ \left(\mathbf{P}^e \mathbf{H}^\top \right) \left[\mathbf{R} + \rho_{yy} \circ \left(\mathbf{H}\mathbf{P}^e \mathbf{H}^\top \right) \right]^{-1}. \quad (9)$$

The local ensemble square root Kalman filter (LEnSRF)

- Perturbation update of the global EnSRF (in state space by definition):

$$\mathbf{X}_a = \mathbf{T}\mathbf{X}_f \quad \text{with} \quad \mathbf{T}_x = \left(\mathbf{I}_x + \mathbf{X}_f \mathbf{X}_f^\top \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{-\frac{1}{2}}.$$

- Covariance localisation: $\mathbf{X}_f \mathbf{X}_f^\top \longrightarrow \mathbf{B} = \rho \circ (\mathbf{X}_f \mathbf{X}_f^\top)$,

$$\mathbf{X}_a = \mathbf{T}\mathbf{X}_f \quad \text{with} \quad \mathbf{T}_x = \left(\mathbf{I}_x + \mathbf{B} \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{-\frac{1}{2}}.$$

The LEnSRF: mode expansion

- ▶ The LEnSRF requires the inverse square root of an $N_x \times N_x$ matrix. Too costly!
- ▶ We wish to make a mode expansion $\mathbf{B} = \rho \circ (\mathbf{X}_f \mathbf{X}_f^\top) \approx \mathbf{X}_r \mathbf{X}_r^\top$, where $\mathbf{X}_r \in \mathbb{R}^{N_x \times N_r}$. If we can do so, we will be able to make a perturbation à la ETKF in the expansion mode subspace rather than in the ensemble subspace.
- ▶ For high-dimensional chaotic models, we would typically have: $N_e \ll N_r \ll N_x$.
- ▶ **The mathematical problem**
Given the matrix $\mathbf{B} = \rho \circ (\mathbf{X}_f \mathbf{X}_f^\top)$, we want to construct a matrix $\mathbf{X}_r \in \mathbb{R}^{N_x \times N_r}$ such that

$$\mathbf{X}_r \mathbf{X}_r^\top \approx \mathbf{B} \quad \text{and} \quad \mathbf{X}_r \mathbf{1} = \mathbf{0}.$$

The LEnSRF: modulation

- ▶ Suppose that there is a matrix \mathbf{W} with N_r columns such that $\rho \approx \mathbf{W}\mathbf{W}^\top$.
- ▶ We define the **modulation product** of \mathbf{W} and \mathbf{X}_f as the matrix with $N_r N_e$ columns:

$$[\mathbf{W}\Delta\mathbf{X}_f]^{jN_e+i} = [\mathbf{W}]_n^j [\mathbf{X}_f]_n^i.$$

This is a mix between a Schur product (for the state variable index n) and a tensor product (for the ensemble indices i and j) [Buehner 2005].

The matrix $\mathbf{X}_r = \mathbf{W}\Delta\mathbf{X}_f$ is a solution with $N_r = N_m N_e$ columns to the problem

$$\mathbf{X}_m \mathbf{X}_m^\top \approx \mathbf{B} \quad \text{and} \quad \mathbf{X}_m \mathbf{1} = \mathbf{0}.$$

- ▶ The modulation product is based on a factorisation property shown by [Lorenc 2003] and is currently used for covariance localisation [Bishop, Whitaker, et al. 2017], including in operational centres [Arbogast et al. 2017].

The LEnSRF: the randomised SVD approach

- ▶ Direct mode expansion of $\rho \circ \mathbf{P}^e$: a singular value decomposition (SVD) is unfeasible!
 - ▶ The **randomised SVD** is an alternative to the **Lanczos** method.
 - (i) It defines a **reduced random subspace** in the column-space of $\rho \circ \mathbf{P}^e$. This subspace is generated by the application of $\rho \circ \mathbf{P}^e$ on random vectors \mathbf{v} : $\rho \circ \mathbf{P}^e \cdot \mathbf{v}$.
 - (ii) A regular svd is then performed in the generated subspace.
 - ▶ Rigorous probabilistic bounds can be obtained on the SVD, given the number of desired modes [Halko et al. 2011].
 - ▶ **Critical advantage**: the application of $\rho \circ \mathbf{P}^e$ on the random vectors \mathbf{v} are independent and are hence carried out **in parallel**.
 - ▶ It was applied to the local EnSRF in [Farchi and Bocquet 2019] .
- Much more on randomised SVD in Omar Ghattas' lectures next week!

The LEnSRF: mode expansion

- Let us assume a mode expansion $\mathbf{B} = \rho \circ (\mathbf{X}_f \mathbf{X}_f^\top) \approx \mathbf{X}_r \mathbf{X}_r^\top$.

$$\mathbf{X}_a \approx \mathbf{T} \mathbf{X}_f \quad \text{with} \quad \mathbf{T}_x = \left(\mathbf{I}_x + \mathbf{X}_r \mathbf{X}_r^\top \mathbf{H}^\top \mathbf{R}^{-1} \mathbf{H} \right)^{-\frac{1}{2}}.$$

Let us use the last algebraic identity and obtain

$$\mathbf{X}_a = \mathbf{T}_r \mathbf{X}_f \quad \text{with} \quad \mathbf{T}_r = \mathbf{I}_x - \mathbf{X}_r \left(\mathbf{I}_r + \mathbf{Y}_r^\top \mathbf{R}^{-1} \mathbf{Y}_r + \left[\mathbf{I}_r + \mathbf{Y}_r^\top \mathbf{R}^{-1} \mathbf{Y}_r \right]^{\frac{1}{2}} \right)^{-1} \mathbf{Y}_r^\top \mathbf{R}^{-1} \mathbf{H}.$$

Now, the algebra is performed in the reduced/mode subspace. It has been proposed in [Bocquet 2016] and later called the Gain Form of the ensemble transform Kalman filter in [Bishop, Whitaker, et al. 2017].

- An approximation which avoids the square root, similar to the DEnKF, is

$$\mathbf{X}_a \approx \mathbf{X}_f - \frac{1}{2} \left(\mathbf{I}_r + \mathbf{Y}_r^\top \mathbf{R}^{-1} \mathbf{Y}_r \right)^{-1} \mathbf{Y}_r^\top \mathbf{R}^{-1} \mathbf{H} \mathbf{X}_f.$$

A multilayer extension of the L96 model

- We introduce the **mL96 model**, which consists of $P_z = 32$ coupled layers of the L96 model with $P_h = 40$ variables:

$$\begin{aligned} \frac{dx_{(z,h)}}{dt} = & \left(x_{(z,h+1)} - x_{(z,h-2)} \right) x_{(z,h-1)} - x_{(z,h)} + F_z \\ & + \underbrace{\delta_{\{z>0\}} \left(x_{(z-1,h)} - x_{(z,h)} \right)}_{\text{Coupling from below}} \\ & + \underbrace{\delta_{\{z<P_z\}} \left(x_{(z+1,h)} - x_{(z,h)} \right)}_{\text{Coupling from above}}. \end{aligned}$$

- The forcing term linearly (and realistically) decreases from $F_1 = 8$ to $F_{32} = 4$.

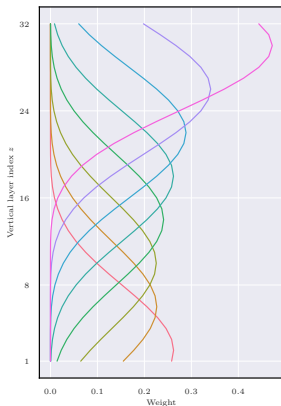
Satellite observations for the mL96 model

- ▶ Each column is observed independently via:

$$y_{c,h} = \sum_{z=1}^{P_z} [\mathbf{\Omega}]_{c,z} x_{z,h} + v_{c,h}, \quad v_{c,h} \sim \mathcal{N}(0,1),$$

where $\mathbf{\Omega}$ is a weighting matrix with $N_c = 8$ channels that is designed to mimic **satellite radiances**.

- ▶ The 8×40 observations are available every $\Delta t = 0.05$.
- ▶ The runs are $10^4 \Delta t$ long.
- ▶ All algorithms use an ensemble of $N_e = 8$ members.

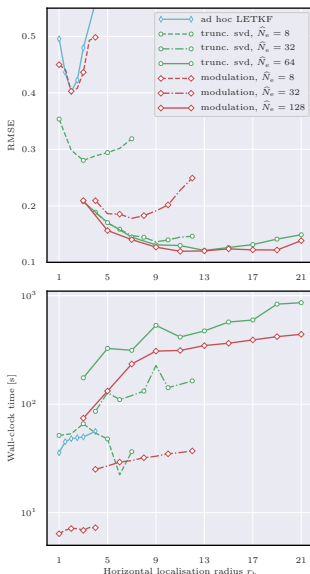


Covariance localisation (with augmented ensembles) is used **only in the vertical direction**. Domain localisation (LETKF-like) is used in the horizontal direction.

Results with the mL96 model

- ▶ Using covariance localisation in the vertical direction yields better RMSE scores than the LETKF.
- ▶ The **modulation** method requires a larger augmented ensemble size to yield similar RMSE scores as the **randomised SVD** method.
- ▶ Both methods benefit from the parallelisation of the local analyses, but the parallelisation potential of the **randomised SVD** method is not fully exploited because of a limited number of threads.

[Farchi and Bocquet 2019]



References I

- [1] B. D. O. Anderson and J. B. Moore. *Optimal Filtering*. Englewood Cliffs, New Jersey: Prentice-Hall, Inc, 1979, p. 357.
- [2] J. L. Anderson. "An ensemble adjustment Kalman filter for data assimilation". In: *Mon. Wea. Rev.* 129 (2001), pp. 2884–2903.
- [3] A. Andrews. "A square root formulation of the Kalman covariance equations". In: *AIAA J.* 6 (1968), pp. 1165–1166.
- [4] E. Arbogast, G. Desroziers, and L. Berre. "A parallel implementation of a 4DEnVar ensemble". In: *Q. J. R. Meteorol. Soc.* 143 (2017), pp. 2073–2083.
- [5] M. Asch, M. Bocquet, and M. Nodet. *Data Assimilation: Methods, Algorithms, and Applications*. Fundamentals of Algorithms. SIAM, Philadelphia, 2016, p. 324.
- [6] C. H. Bishop, B. J. Etherton, and S. J. Majumdar. "Adaptive Sampling with the Ensemble Transform Kalman Filter. Part I: Theoretical Aspects". In: *Mon. Wea. Rev.* 129 (2001), pp. 420–436.
- [7] C. H. Bishop, J. S. Whitaker, and L. Lei. "Gain form of the Ensemble Transform Kalman Filter and its relevance to satellite data assimilation with model space ensemble covariance localization". In: *Mon. Wea. Rev.* 145 (2017), pp. 4575–4592.
- [8] M. Bocquet. "Localization and the iterative ensemble Kalman smoother". In: *Q. J. R. Meteorol. Soc.* 142 (2016), pp. 1075–1089.
- [9] M. Bocquet and A. Farchi. "On the consistency of the perturbation update of local ensemble square root Kalman filters". In: *Tellus A* 71 (2019), pp. 1–21.
- [10] M. Bocquet and P. Sakov. "An iterative ensemble Kalman smoother". In: *Q. J. R. Meteorol. Soc.* 140 (2014), pp. 1521–1535.
- [11] M. Buehner. "Ensemble-derived stationary and flow-dependent background-error covariances: Evaluation in a quasi-operational NWP setting". In: *Q. J. R. Meteorol. Soc.* 131 (2005), pp. 1013–1043.
- [12] G. Burgers, P. J. van Leeuwen, and G. Evensen. "Analysis scheme in the ensemble Kalman filter". In: *Mon. Wea. Rev.* 126 (1998), pp. 1719–1724.
- [13] A. Carrassi et al. "Data Assimilation in the Geosciences: An overview on methods, issues, and perspectives". In: *WIREs Climate Change* 9 (2018), e535.
- [14] S. E. Cohn, N. S. Sivakumaran, and R. Todling. "A Fixed-Lag Kalman Smoother for Retrospective Data Assimilation". In: *Mon. Wea. Rev.* 122 (1994), pp. 2838–2867.
- [15] E. Cosme et al. "Smoothing problems in a Bayesian framework and their linear Gaussian solutions". In: *Mon. Wea. Rev.* 140 (2012), pp. 683–695.
- [16] R. Daley. *Atmospheric Data Analysis*. Cambridge University Press, New-York, 1991, p. 472.

References II

- [17] M. El Ghararni. "Enhanced Adaptive Inflation Algorithm for Ensemble Filters". In: *Mon. Wea. Rev.* 146 (2018), pp. 623–640.
- [18] G. Evensen. *Data Assimilation: The Ensemble Kalman Filter. Second.* Springer-Verlag Berlin Heidelberg, 2009, p. 307.
- [19] G. Evensen. "Sequential data assimilation with a nonlinear quasi-geostrophic model using Monte Carlo methods to forecast error statistics". In: *J. Geophys. Res.* 99 (1994), pp. 10143–10162.
- [20] G. Evensen. "The Ensemble Kalman Filter: Theoretical Formulation and Practical Implementation". In: *Ocean Dynamics* 53 (2003), pp. 343–367.
- [21] G. Evensen and P. J. van Leeuwen. "An Ensemble Kalman Smoother for Nonlinear Dynamics". In: *Mon. Wea. Rev.* 128 (2000), pp. 1852–1867.
- [22] A. Farchi and M. Bocquet. "On the efficiency of covariance localisation of the ensemble Kalman filter using augmented ensembles". In: *Front. Appl. Math. Stat.* 5 (2019), p. 3.
- [23] S. J. Fletcher. *Data assimilation for the geosciences: From theory to application.* Elsevier, 2017.
- [24] M. Ghil and P. Malanotte-Rizzoli. "Data assimilation in meteorological and oceanography". In: *Advanc. in Geophys.* 33 (1991), pp. 141–266.
- [25] N. Halko, P.-G. Martinsson, and J. A. Tropp. "Finding structure with randomness: Probabilistic algorithms for constructing approximate matrix decompositions". In: *SIAM review* 53 (2011), pp. 217–288.
- [26] T. M. Hamill, J. S. Whitaker, and C. Snyder. "Distance-dependent filtering of background error covariance estimates in an ensemble Kalman filter". In: *Mon. Wea. Rev.* 129 (2001), pp. 2776–2790.
- [27] N. J. Higham. *Functions of matrices: theory and computation.* Vol. 104. Siam, 2008, p. 450.
- [28] P. L. Houtekamer and H. L. Mitchell. "A sequential ensemble Kalman filter for atmospheric data assimilation". In: *Mon. Wea. Rev.* 129 (2001), pp. 123–137.
- [29] P. L. Houtekamer and H. L. Mitchell. "Data assimilation using an ensemble Kalman filter technique". In: *Mon. Wea. Rev.* 126 (1998), pp. 796–811.
- [30] B. R. Hunt, E. J. Kostelich, and I. Szunyogh. "Efficient data assimilation for spatiotemporal chaos: A local ensemble transform Kalman filter". In: *Physica D* 230 (2007), pp. 112–126.
- [31] E. Kalnay. *Atmospheric Modeling, Data Assimilation and Predictability.* Cambridge University Press, Cambridge, 2002, p. 357.
- [32] W. G. Lawson and J. A. Hansen. "Implications of Stochastic and Deterministic Filters as Ensemble-Based Data Assimilation Methods in Varying Regimes of Error Growth". In: *Mon. Wea. Rev.* 132 (2004), pp. 1966–1981.

References III

- [33] D. M. Livings, S. L. Dance, and N. K. Nichols. "Unbiased ensemble square root filters". In: *Physica D* 237 (2008), pp. 1021–1028.
- [34] A. C. Lorenc. "The potential of the ensemble Kalman filter for NWP - a comparison with 4D-Var". In: *Q. J. R. Meteorol. Soc.* 129 (2003), pp. 3183–3203.
- [35] E. N. Lorenz and K. A. Emanuel. "Optimal sites for supplementary weather observations: simulation with a small model". In: *J. Atmos. Sci.* 55 (1998), pp. 399–414.
- [36] E. Ott et al. "A local ensemble Kalman filter for atmospheric data assimilation". In: *Tellus A* 56 (2004), pp. 415–428.
- [37] P. N. Raanes, M. Bocquet, and A. Carrassi. "Adaptive covariance inflation in the ensemble Kalman filter by Gaussian scale mixtures". In: *Q. J. R. Meteorol. Soc.* 145 (2019), pp. 53–75. eprint: [arXiv:1801.08474](https://arxiv.org/abs/1801.08474).
- [38] S. Reich and C. Cotter. *Probabilistic Forecasting and Bayesian Data Assimilation*. Cambridge University Press, 2015, p. 306.
- [39] P. Sakov and L. Bertino. "Relation between two common localisation methods for the EnKF". In: *Comput. Geosci.* 15 (2011), pp. 225–237.
- [40] P. Sakov and M. Bocquet. "Asynchronous data assimilation with the EnKF in presence of additive model error". In: *Tellus A* 70 (2018), p. 1414545.
- [41] P. Sakov, G. Evensen, and L. Bertino. "Asynchronous data assimilation with the EnKF". In: *Tellus A* 62 (2010), pp. 24–29.
- [42] P. Sakov and P. R. Oke. "A deterministic formulation of the ensemble Kalman filter: an alternative to ensemble square root filters". In: *Tellus A* 60 (2008), pp. 361–371.
- [43] P. Sakov and P. R. Oke. "Implications of the Form of the Ensemble Transformation in the Ensemble Square Root Filters". In: *Mon. Wea. Rev.* 136 (2008), pp. 1042–1053.
- [44] J. S. Whitaker and T. M. Hamill. "Ensemble Data Assimilation without Perturbed Observations". In: *Mon. Wea. Rev.* 130 (2002), pp. 1913–1924.